# Predicting the Effect of Business Births and Deaths on the Current Employment Statistics Survey: Using Sample Information to Minimize Coverage Error 

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#### Abstract

The Current Employment Statistics (CES) program uses a model to account for the bias in monthly payroll employment estimates arising from establishment births and deaths that fall outside the survey frame. Actual birth-death values derived from administrative counts are available with a substantial lag but must be predicted for the current month's estimation. The existing CES net birth-death model relies on the bias from business births and deaths to follow a consistent, seasonal pattern, characterized by an ARIMA process. This has broken down during extreme changes in the labor market necessitating interventions in the model during the COVID-19 recession and recovery. Previous research showed the ARIMA models can be improved by including covariates available coincident with the survey. This paper explores several modelling frameworks that use information from the CES survey to predict birth-death values, substantially outperforming previously examined models over the period from 2007-2020, covering the Great Recession through the pandemic. Forecast combination techniques are also examined and compared with predictions from the individual frameworks.


Key Words: Nowcasting, birth-death model, forecast combination

## 1. Introduction

The Current Employment Statistics (CES) program at the U.S. Bureau of Labor Statistics (BLS) produces nonfarm payroll employment, hours, and earnings data each month at the national level, for all 50 states, the District of Columbia, Puerto Rico, and about 450 metropolitan areas and divisions. National level data are typically released on the first Friday following the end of the reference month, representing one of the timeliest Principal Federal Economic Indicators, and are closely watched by policy makers, markets, and others as a coincident measure of the U.S. economy.

CES data are produced using a monthly establishment survey but include an adjustment-the net birthdeath forecast-due to the survey's inability to capture new businesses ("births") and difficulty capturing closures ("deaths") in real time. BLS can calculate the "actual birth-death value"-the adjustment needed to offset this coverage error-but only at a lag of 10-12 months behind the current reference period. BLS uses these actual birth-death values as input to a seasonal ARIMA time series model forward to the current month. This captures the seasonal pattern in birth-death values but does a poor job reflecting business cycle variation since it does not include any real-time information.

Using concurrent information to predict the actual birth-death values can be considered a "nowcasting" problem. Previous work by Battista (2013) showed that covariates derived from the CES survey or other

[^0]timely economic indicators added to the ARIMA models would have improved predictive accuracy during the Great Recession (2007-09). CES used a modified version of this method during the steep downturn of the COVID-19 pandemic, but found its impact to be insufficient, and made temporary adjustments to the survey estimator to capture the effect of short-term business closures and reopenings.

We find that predictions of the actual birth-death values can be substantially improved by using information in the CES survey without making changes to the survey estimator. We simulate 44 individual prediction approaches over the period 2007-2020, many of which give good out-of-sample results, and which typically encompass the true birth-death values. This holds over all phases of the business cycle including the Great Recession, the ensuing recovery/expansion, and the extreme events of the pandemic. We would not have known the best ex post approach ex-ante and all likely suffer from some degree of model misspecification. Therefore, we consider forecast combination and find relatively simple techniques to work well.

## 2. Coverage Error from Business Births and Deaths

The monthly CES survey contains about 122,000 businesses and government agencies representing approximately 666,000 establishments reporting positive employment in the current and prior month. CES cannot capture births since they are not on the survey frame, the BLS Longitudinal Database (LDB), which is derived from the universe of Unemployment Insurance (UI) tax records, available at a lag of several months. Deaths are likewise difficult to capture since permanent closures generally result in nonresponse, and it is not practicable to determine the status of all nonrespondents during the short timeline needed for production. Some closures are reported-often temporary ones, or closures of establishments that were part of a multiple worksite employer.

The CES Net Birth-Death model was introduced as part of a probability-based redesign of the CES survey undertaken in the 1990s (Mueller 2006), replacing a bias-adjustment model that accounted for births, deaths, and other sources of bias in the previous quota sample. Nonrespondents and those businesses that report a closure are implicitly imputed with the sample weighted link relative (equation 1), which CES uses to estimate relative employment growth among continuing businesses, calculated within each estimating cell (detailed industry ${ }^{2}$ ) at time $t$ :

Equation 1: Weighted Link Relative

$$
W L R_{t}=\frac{\sum_{i=1}^{n} w_{i} * a e_{i, t}}{\sum_{i=1}^{n} w_{i} * a e_{i, t-1}}
$$

Where ae represents the number of all employees at establishment $\boldsymbol{i}$ in all $\boldsymbol{n}$ responding establishments with $\boldsymbol{a} \boldsymbol{e}_{\mathrm{t}}>0$ and $\boldsymbol{a} \boldsymbol{e}_{\boldsymbol{t}-1}>0$, and $\boldsymbol{w}$ is the survey weight associated with each reporter.

This eliminates the need to determine the employment status of nonrespondents and accounts for most business births.

[^1]The historical bias from applying this procedure (i.e., the amount it under- or over-accounts for births) can be determined using data from the LDB: the survey estimator is applied to the population from an initialization period $t=0$, representing a March benchmark month, through $t=24$. The population version of the link relative (also calculated at the estimating cell level) is shown in equation 2 :

Equation 2:Population Link Relative

$$
L R_{t}^{P O P}=\frac{\sum_{i=1}^{N} a e_{i, t}}{\sum_{i=1}^{N} a e_{i, t-1}}
$$

For all $\boldsymbol{N}$ population establishments with $\boldsymbol{a} \boldsymbol{e}_{\boldsymbol{t}}>0, \boldsymbol{a} \boldsymbol{e}_{\boldsymbol{t}-1}>0$, and $\boldsymbol{a} \boldsymbol{e}_{\boldsymbol{t}=0}>0$. The establishment must have positive employment in the initialization month of the frame, $t=0$, which removes any births occurring since the starting period. No weights are applied since all population establishments with positive employment in the current, previous, and initialization months are included.

Next, the Continuous Plus Imputed (CIMP) series from $t=1$ to $t=24$ is derived by applying the population link relative to population employment level from $t=0$ :

Equation 3: Continuous Plus Imputed Series

$$
\text { CIMP }_{t}=A E_{t=0}^{L D B} \prod_{i=1}^{t} L R_{t=i}^{P O P}
$$

The difference between LDB employment and the CIMP represents the cumulative birth-death error from the initialization month. The actual birth-death values are then calculated as the over-the-month change in this residual:

Equation 4: Birth-Death calculated from LDB

$$
B D_{t}=(1-B)\left(A E_{t}^{L D B}-C I M P_{t}\right)
$$

Where $\boldsymbol{B}$ represents the backshift (or lag) operator. ${ }^{3}$
Two full years ( 24 months) of historical birth-death values are calculated from each initialization so that birth-death values can be derived for the appropriate age of the sample used in rotation in CES. (In a given quarter, the sample frame used for some industries is a year older than in other industries.) Several years of BD values calculated from months 1-12 following the initialization period are chained together to form a "Year 1" time series, while values from months 13-24 are chained to form a "Year 2" series. In practice, there is usually little difference between the Year 1 and Year 2 values; Year 2 tends to be slightly more positive than Year 1 in each calendar month/industry.

Historical values of $B D_{t}$ must be forecast up to 12 months beyond their end date for use in estimation. Consider that $B D_{t}$ can be characterized as following a seasonal ARIMA process $z_{t}{ }^{4}$ with mean $\beta X_{t}$ :

[^2]$$
B D_{t}=\beta X_{t}+z_{t}
$$

Where $X_{t}$ is a matrix of covariates available at time $t$ and $\beta$ is a vector of coefficients.
Historically, CES has not included covariates in the model (i.e., $X_{t}$ is null ${ }^{5}$ ), and has assumed a seasonal integrated moving average model for $\mathrm{z}_{\mathrm{t}}$ :

Equation 6: Birth-Death as Seasonal IMA

$$
B D_{t}=z_{t}=\Theta\left(B^{12}\right) a_{t} /\left(1-B^{12}\right)
$$

Where $a_{\mathrm{t}}$ are white noise.
Putting everything together, CES employment estimates are created by setting a benchmark population employment level and applying the weighted link relative plus birth-death adjustment for successive months:

Equation 7: CES All Employees Estimator

$$
A E_{t}=A E_{t=0}^{P O P} \prod_{i=1}^{t} W L R_{t=i}+B D_{t=i}
$$

The problems addressed in this paper can be considered as determining what to include in $X_{t}$, estimating $\beta$, and forecasting $z_{t} .{ }^{6}$ We assume that the historical values of $B D_{t}$ derived from the LDB are measured precisely and that our task lies in their prediction. There are sources of nonsampling error in both QCEW and CES that challenge this assumption. However, the birth-death values applied to CES data provide well-centered benchmark revisions when compared to QCEW and improved prediction of the birthdeath values could substantially reduce the size of the benchmark revisions.

### 2.1 Relationship between Birth-Death and CES Sample

Both the actual birth-death values and the CES sample weighed link relative exhibit seasonal, cyclical, and irregular time series characteristics. Figure 1 shows both time series of Year 1 birth-death values and the weighted link relative at the total private level, standardized by the mean and standard deviation of each series prior to 2020. Both series have somewhat similar seasonal patterns, and both have a similar cyclical decline during the Great Recession (Dec. 2007 - Jun. 2009). The degree to which the sample link and birth-death share similar seasonality varies considerably by component industry. They are closely linked in construction and leisure and hospitality, where seasonality is primarily weather-driven and there are many small businesses but have little in common in industries such as education services, with many larger establishments, and seasonality arising from administrative effects

[^3](e.g., the school year). The extreme negative and positive values both series exhibited in 2020 were of comparable magnitude and timing.

Figure 1: Total Private Birth-Death (Year 1) and Weighted Link Relative - Standardized


At a very high level, both birth-death and the sample demonstrate similar cyclical patterns. Table 1 shows annual sums of actual birth-death Year 1 values and the average sample link relative at the total private level. Both showed similar downturns in 2008 and 2009 during the Great Recession, although the sample recovered earlier than the birth-death values in 2010-2011.

Table 1: Year 1 Birth-Death (annual total, thousands) and Weighted Link Relative (annual average)

|  | 2007 | 2008 | 2009 | 2010 | 2011 | 2012 | 2013 | 2014 | 2015 | 2016 | 2017 | 2018 | 2019 | 2020 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| B-D (thousands) | 755 | 278 | 18 | 474 | 606 | 964 | 803 | 880 | 993 | 823 | 908 | 911 | 879 | 75 |
| Weighted Link Relative | 1.000 | 0.997 | 0.996 | 1.001 | 1.001 | 1.001 | 1.001 | 1.002 | 1.001 | 1.001 | 1.001 | 1.001 | 1.001 | 0.995 |

Note that the weighted link relative never accounts for more than a couple tenths of a percentage point of annual growth, while actual net birth-death can account for close to 1 million jobs, or about 7 tenths of a percent. (Total private annual employment ranged between about 115 million and 130 million over this period.) This aligns with Haltiwanger (2013), which demonstrated that most net job creation
stemmed from growth in new and young businesses. Much of this growth happens in businesses too new for the CES sample to capture. However, information in the CES sample can help predict net job growth among new businesses and business closures.

## 3. Predicting Actual Birth-Death Values: Simulations

We simulated birth-death predictions for April 2007 - December 2020 for 3 baseline models, 44 individual forecasting models, and 6 types of forecast combinations. We accounted for rotation of Year 1/Year 2 values according to CES sample rotation.

We considered five categories of approaches for predicting birth-death values:

1. Baseline ARIMA/RegARIMA
2. Extended RegARIMA
3. Partially cross-sectional
4. Recurrent Neural Networks
5. Forecast combinations

Each will be broadly described in this section.
All models made use of a subset of shared set of covariates based on the CES sample. These included:

- The sample link relative $\left(W L R_{t}\right)$. In some cases, we used the link relative at a higher level of industry detail to model at the detailed level. In many cases we took the natural log of the link relative, which often gave better diagnostics, but gave similar results in most cases.
- A measure of reported drops-to and returns-from zero employment in the CES sample.
- Seasonal dummy variables.
- Other information about the CES sample, such as the proportion of the sample below a certain employment size threshold.
- Detailed information about the CES sample including the amount of employment by establishment size increasing or decreasing within a certain range.

We produced our models at up to 3 different levels of industry aggregation, each of which we considered to be a separate model. CES needs basic industry (up to 6-digit NAICS level) detailed birthdeath values in production. However, we also ran simulations at the 3-digit NAICS and super sector (generally 2-digit) level, which can serve as a control for the sum of basic level predictions. (This sort of reconciliation is currently done to produce birth-death forecasts at the state and metro area level.) For our results, we sum to the super sector or total private and evaluate at that level.

### 3.1 Baseline ARIMA/RegARIMA

We replicated the seasonal ARIMA models that CES has used since the introduction of the net birthdeath models in 2003. We used automatic outlier selection in SAS PROC X13 (SAS 2018) instead of outliers selected during a manual review by BLS analysts, so the forecasts in our replication did not perfectly match those created in production, but the overall results were similar enough to use as a baseline.

We also replicated the approach from Battista (2013) and a variant that was used in production in 2020. This method adds the sample links for each super sector ${ }^{7}$ to the RegARIMA model estimated at the basic level. When the model was applied in 2020, a considerable number of series exhibited negative coefficients for the relationship between the sample link and the net birth-death value-two series that presumably should move together rather than in opposition. In those cases, and others where modeling diagnostics were unacceptable, the seasonal difference was removed from the ARIMA part of the model and seasonal dummy variables were added. This usually resulted in predictions very similar to the baseline ARIMA, so we substituted with those forecast results under the same conditions. With all simulations, in the rare cases where models failed to converge, we substituted with the baseline ARIMA.

### 3.2 Extended RegARIMA

We considered it preferable to use established, easily interpretable models. To that end, we extended the use of RegARIMA models. We lengthened our input series as far as possible and used the sample link calculated for the same domain we modeled birth-death. Some variants included a measure of reported drops-to and returns-from zero as a covariate or added lags of the covariates.

### 3.3 Partially Cross-Sectional

Many of the methods we explored use the following 2-step approach. In the first step, we ignored the time series nature of the data and estimated a cross sectional regression relating birth-death to the covariates. The residuals from this first step were autocorrelated, so as a second step, we forecasted the residuals with a seasonal ARIMA model.

The primary motivation for this approach follows from the fact that (outside of the COVID-19 pandemic) most of the variation in birth-death values and the CES sample is seasonal, that the seasonality is likely cointegrated, and that this simple approach gives good out-of-sample results. Our intuition follows Barsky and Miron (1989), who found close links between seasonality and the business cycle. We think that in many industries, the same causes drive seasonal and business cycle fluctuations in both the sample and birth-death. Seasonal differencing or fitting seasonal dummies may result in model misspecification if the series are cointegrated. In future work, we would like to thoroughly investigate the possible seasonal cointegrating relationships.

The secondary motivation for this approach was that, in many detailed series, sampling error and other noise obscures the relationship between birth-death and the sample link. Pooling information across all series was easily handled in this approach but could also be done with multivariate time series models.

We tried many slightly different approaches within this partially cross-sectional framework. At its simplest, we estimated a univariate linear regression with the sample link. Variations included adding sets of seasonal dummy variables. We had some success fitting an initial regression, finding some structure in the residuals by calendar month, then constructing clustered seasonal dummies for the final stage 1 regression ${ }^{8}$. (E.g., we may form one intercept for Jan., Apr., and Jun., and another for the other 9 months.) This seemed to strike a balance between setting a full set of seasonal dummies and ignoring seasonality all together. We also had some success in the pooled approach constructing many covariates

[^4]out of the CES sample based on establishment size. The intuition there was that the amount of growth or loss among smaller worksites would be closely linked to births and deaths, but the relationship would attenuate among larger businesses.

### 3.4 Recurrent Neural Networks

Neural networks have been explored extensively recently in many areas including time series forecasting. One such python package created for this purpose is GlutonTS (Alexandrov 2019). We used the DeepAR (Deep Autoregressive) model from the GluonTS in our research. A DeepAR model is an autoregressive recurrent neural network that is trained on multiple series at once to learn a global structure that allows probabilistic forecasts (Salinas 2020). The model allows for the use of covariates and multi-horizon forecasts. We used the same covariates as the other methods in this model.

The motivation for this method is to experiment with newer machine learning methods for time series forecasting at BLS. We applied this method with the default settings with moderate success, however more experimentation with these models will be needed.

### 3.5 Forecast Combination

Combinations of economic forecasts have long been found to perform well compared with individual forecasts, dating back to at least Bates and Granger (1969), who proposed that information available to expert forecasters may not be available to the individual forecaster. Forecast combination can also serve to make predictions more robust to model misspecification bias—our models are at best rough approximations to the true data generating process-which serves as our motivation for combining forecasts to predict birth-death (Timmermann 2006).

We considered two different sets of candidate forecasts for combination:

- $\quad S_{1}$ : All forecasting models, excluding baseline approaches
- $S_{2}$ : Partially cross-sectional forecasts without full seasonal dummies ( $\mathrm{S}_{2} \subseteq \mathrm{~S}_{1}$ )

The reason for the second set is that these approaches did not seem to suffer from the same biases as other approaches but individually could be noisy and flawed.

With each candidate set we investigated three forecast combination approaches for our birth-death simulations. The first forecast combination approach is a simple average of the set of candidate forecasts ( $f_{i} \in S_{j}$ ):

Equation 7: Simple Average Forecast Combination

$$
\widehat{B D}_{t}^{j}=N^{-1} \sum_{i}^{N} f_{i, t}
$$

Very often simple averages empirically outperform forecast combinations that attempt to find optimal weights, which has become known as the "forecast combination puzzle" (Steel 2020).

The second approach is from Bates and Granger (1969), which weights based on the error variance of the individual forecasts:

$$
\widehat{B D}_{t}^{j}=\sum_{i}^{N} w_{i}^{B G} f_{i, t}, w_{i}^{B G}=\frac{\hat{\sigma}_{i}^{-2}}{\sum_{j=1}^{N} \hat{\sigma}_{j}^{-2}}
$$

We estimated the variances based on prior out-of-sample forecast error. (Our first year of simulations used equal weights.)

The third approach uses a sophisticated subset averaging developed by Diebold and Shin (2019), who found the best approach in an empirical study was to incorporate principles of selection and regularization, discard most forecasts, and use equal weights for the rest. Their approach searches for the "best average" combination of $\mathrm{N} \leq \mathrm{K}$ candidate forecasts based on prior out-of-sample results. This can be very computationally expensive ${ }^{9}$ and we set $K=5$.

## 4. Results

We evaluated two different out-of-sample error measures. The first (M1) was a straightforward root mean square error measure of the predicted birth-death values evaluated at the super sector level, based on an information set containing actual birth-death values 10-12 months behind the end of the nowcasting horizon, for each method (j). E.g., we evaluated Jan.-Mar. 2017 predictions with birth-death actuals through Mar. 2016 and covariates through Mar. 2017. This mirrors the information CES has when producing the initial monthly employment estimates. Results are presented by year as well as across all years since data users care about performance across the business cycle.

Equation 9: RMSE Monthly Birth-Death

$$
M 1_{j}=\sqrt{N_{t}^{-1} \sum_{t=1}^{N_{t}}\left(\widehat{B D}_{t, j}-B D_{t}\right)^{2}}
$$

The second measure (M2) relates to the total error in our post-benchmark re-estimates. After setting a March benchmark level, CES updates birth-death values for the following 9 months in addition to producing initial forecasts for Jan.-Mar. of the following year. We sum the 12 months of predictions and actuals to calculate a root mean square error measure evaluated at the total private level. Years are designated by the year of the last month being predicted. (E.g., "2015" covers errors in the summed birth-death predictions over Apr. 2014 - Mar. 2015.)

Equation 10: RMSE 12-Mo. Post-Benchmark

$$
M 2_{j}=\sqrt{N_{\text {year }}^{-1} \sum_{\text {year }=1}^{N_{\text {year }}}\left(\sum_{m=1}^{12}\left(\widehat{B D}_{t, j}-B D_{t}\right)\right)^{2}}
$$

Summing by year across all industries is done because errors around business cycle turning points have often been correlated across industries for several months in a row. For example, the production birthdeath forecasts covering the worst of the Great Recession were too high broadly across industries over

[^5]several months since few new businesses formed during a financial crisis while many closed; forecasts were then too low during the ensuing recover/expansion.

Due to the large number of forecasting models, we present the baseline models, the forecast combinations, and results for the five best, median, and worst-performing individual model (evaluated across all years). Results for all individual methods can be found in the appendix.

The version of Battista (2013) with modifications made to the RegARIMA models in 2020 was the best baseline approach, which we compared against the alternatives we investigated.

Under Metric 1 (Table 2), our best ex post individual forecast (N28)—which included zeros and clustered seasonal intercepts in the regression-cut the RMSE nearly in half compared to that baseline. Most of this gain was achieved in 2020. Nearly all individual methods outperformed the best baseline. The simple average and Bates-Granger forecast combinations using the subset of methods $\left(\mathrm{S}_{2}\right)$ performed nearly as well as the best ex post individual. These methods did not perform quite as well as the best ex post forecast in 2020 but tended to perform somewhat better in other years. The combinations performed well more consistently than any individual approach. Simple averaging and Bates-Granger averaging using the full set of individuals $\left(S_{1}\right)$ performed only slightly worse than using the subset. The Diebold-Shin approach did not perform as well as the other combinations, driven by underperformance in 2020. With the extreme volatility that year, a strategy to discard most forecasts and average the rest did not work as well.

Table 2 - RMSE, Super Sector Monthly Birth-Death, Thousands (M1)

|  | Method | Overall | 2008 | '09 | '10 | '11 | '12 | '13 | '14 | '15 | '16 | '17 | '18 | '19 | '20 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| ¢ | Production (ARIMA) | 27.1 | 6.9 | 7.7 | 6.2 | 5.4 | 8.2 | 4.3 | 4.4 | 6.2 | 7.2 | 6 | 4.9 | 5 | 95.3 |
|  | RegARIMA - Battista 2013 | 23.2 | 5.7 | 6.7 | 5.8 | 5.8 | 8.2 | 4.3 | 4.5 | 6.3 | 7.3 | 5.8 | 5.1 | 4.9 | 81.1 |
|  | RegARIMA - Battista 2013 (modified) | 20.1 | 5.3 | 6.4 | 5.8 | 5.7 | 8.2 | 4.3 | 4.5 | 5.9 | 7.3 | 5.9 | 5.1 | 4.8 | 69.6 |
| $\begin{aligned} & \text { n } \\ & .0 \\ & 0 \\ & 0 \\ & 0 \\ & 0 \\ & \\ & 0 \\ & 0 \end{aligned}$ | Simple Avg. (S1) | 12.8 | 5.2 | 6.1 | 6.1 | 6.2 | 8 | 4.4 | 4.3 | 6.2 | 7.1 | 5.7 | 5.1 | 5.1 | 41.3 |
|  | Simple Avg. (S2) | 10.9 | 4.7 | 6.1 | 6.6 | 6.8 | 8.1 | 4.6 | 4.4 | 6.2 | 7.1 | 5.7 | 5.4 | 5.1 | 33.2 |
|  | Bates-Granger (S1) | 12.5 | 5.2 | 6.2 | 6.1 | 6.3 | 8 | 4.4 | 4.3 | 6.2 | 7.1 | 5.7 | 5.2 | 5.1 | 40.2 |
|  | Bates-Granger (S2) | 10.9 | 4.7 | 6.2 | 6.7 | 7 | 8.1 | 4.7 | 4.4 | 6.2 | 7.2 | 5.8 | 5.4 | 5.1 | 33.4 |
|  | Diebold-Chen (S1) | 15.4 | 5.2 | 6.1 | 6.4 | 6.5 | 8.3 | 4.6 | 4.4 | 6.4 | 7.1 | 5.7 | 5.1 | 5 | 51.6 |
|  | Diebold-Chen (S2) | 13.4 | 4.7 | 6.2 | 6 | 6.3 | 8.3 | 4.8 | 4.5 | 6.3 | 7.1 | 5.7 | 5.2 | 5.1 | 43.8 |
|  | Rank 1 Individual (N28) | 10.7 | 4.7 | 6.8 | 9.3 | 8.8 | 8.1 | 5.2 | 4.6 | 6.7 | 7.3 | 6.1 | 5.6 | 5.5 | 30.7 |
|  | Rank 2 Individual (N31) | 11 | 5 | 7.4 | 9.9 | 9.7 | 8.2 | 5.9 | 4.8 | 7.3 | 7.2 | 6.2 | 5.6 | 5.5 | 31.2 |
|  | Rank 3 Individual (N08) | 11.1 | 4.7 | 6.6 | 6.2 | 6.8 | 8.1 | 4.5 | 4.5 | 6.4 | 7.2 | 6.3 | 5.3 | 5.5 | 34.1 |
|  | Rank 4 Individual (N11) | 11.4 | 4.7 | 6.8 | 6.2 | 6.9 | 8.1 | 4.5 | 4.5 | 6.4 | 7.3 | 6.2 | 5.4 | 5.5 | 35.2 |
|  | Rank 5 Individual (N09) | 11.8 | 4.7 | 6.4 | 6.1 | 6.7 | 8 | 4.5 | 4.5 | 6.3 | 7.2 | 6.1 | 5.4 | 5.5 | 37 |
|  | Rank 22 (Median) Individual (B28) | 16.2 | 5.8 | 6.6 | 9.2 | 7.9 | 8.4 | 5.3 | 4.6 | 7 | 7.6 | 5.7 | 5.7 | 5.3 | 53.5 |
|  | Rank 44 (Worst) Individual (RNN) | 25.1 | 8.6 | 10.4 | 7.9 | 6.9 | 8.1 | 6.8 | 5.6 | 7.5 | 7.8 | 7.7 | 7 | 8.9 | 86.4 |

Table 3: RMSE, Total Private Post-Benchmark 12-Month Summed Birth-Death, Thousands (M2)

|  | Method | Overall | 2008 | ‘09 | '10 | '11 | '12 | '13 | '14 | '15 | '16 | '17 | '18 | '19 | '20 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Production (ARIMA) | 293 | 179 | 723 | 345 | 157 | 433 | 251 | 222 | 108 | 115 | 149 | 38 | 65 | 257 |
|  | RegARIMA - Battista 2013 | 225 | 104 | 486 | 355 | 150 | 297 | 176 | 203 | 72 | 101 | 168 | 55 | 54 | 225 |
|  | RegARIMA - Battista 2013 (modified) | 214 | 76 | 416 | 373 | 136 | 301 | 187 | 209 | 62 | 101 | 180 | 63 | 56 | 197 |
|  | Simple Avg. (S1) | 201 | 142 | 398 | 348 | 129 | 252 | 123 | 171 | 67 | 95 | 188 | 58 | 73 | 206 |
|  | Simple Avg. (S2) | 175 | 106 | 238 | 354 | 210 | 204 | 103 | 151 | 49 | 82 | 191 | 65 | 83 | 176 |
|  | Bates-Granger (S1) | 199 | 142 | 395 | 342 | 131 | 250 | 124 | 170 | 66 | 98 | 189 | 59 | 73 | 201 |
|  | Bates-Granger (S2) | 174 | 106 | 222 | 347 | 222 | 200 | 105 | 151 | 48 | 85 | 192 | 65 | 83 | 171 |
|  | Diebold-Shin (S1) | 179 | 142 | 226 | 339 | 114 | 259 | 130 | 165 | 68 | 90 | 185 | 39 | 73 | 233 |
|  | Diebold-Shin (S2) | 167 | 106 | 175 | 327 | 123 | 239 | 150 | 150 | 55 | 75 | 176 | 47 | 77 | 220 |
|  | Rank 1 Individual (B60) | 156 | 125 | 105 | 19 | 38 | 278 | 229 | 97 | 61 | 138 | 25 | 83 | 243 | 246 |
|  | Rank 2 Individual (S09) | 161 | 121 | 155 | 309 | 213 | 206 | 88 | 150 | 39 | 101 | 185 | 55 | 105 | 161 |
|  | Rank 3 Individual (S08) | 167 | 98 | 43 | 347 | 291 | 182 | 87 | 141 | 29 | 101 | 189 | 61 | 102 | 152 |
|  | Rank 4 Individual (S15) | 169 | 142 | 167 | 328 | 117 | 244 | 106 | 135 | 22 | 34 | 113 | 9 | 160 | 267 |
|  | Rank 5 Individual (N11) | 172 | 44 | 97 | 415 | 227 | 179 | 58 | 135 | 44 | 103 | 230 | 80 | 46 | 145 |
|  | Rank 22 (Median) Individual (B28) | 219 | 50 | 479 | 389 | 183 | 193 | 142 | 166 | 105 | 124 | 240 | 125 | 22 | 146 |
|  | Rank 44 (Worst) Individual (B27) | 279 | 201 | 685 | 389 | 70 | 326 | 214 | 209 | 131 | 164 | 209 | 92 | 29 | 235 |

Under Metric 2 (Table 3), the best individual methods outperformed the best baseline by over 25 percent, although the best baseline slightly outperformed the median individual method. The forecast combinations using the subset ( $\mathrm{S}_{2}$ ) once again performed comparably to the best ex post individual. In this case, the Diebold-Shin averaging performed best. However, due to the way we set up the simulations, M2 does not capture much of the pandemic period when Diebold-Shin performed poorly.

If we shifted the window for M 2 to sum the 12 months following each December, instead of following the March benchmark, we would better capture the volatility of 2020. In that alternative construction, Diebold-Shin performs comparably to the other combination techniques overall since it does worse in 2020. Overall, our individual methods and forecast combinations tend to show more gain with this alternate window. We must note that, had we evaluated the M2 at the super sector level instead of total private, we would not have seen a substantial gain against the baseline. Both these alternative metrics are reported in the appendix.

Among the individual methods, one of the very simplest partially cross-sectional approaches (designated N11) ranks among the top 5 in both metrics. Its performance is comparable to the better forecast combinations. The most complicated partially cross-sectional approach (B60) performs best on the second metric and has the practical benefit of being modeled at the basic level. The various RegARIMA approaches we tried-adding additional covariates and lags-generally performed worse than the partially cross-sectional models but slightly better than the baselines, at least on Metric 1. The DeepAR approach fared comparably to the baselines. We only tried one approach in this class of models, without contemporaneous information, and we think that it may have more success under different formulations, or as the second stage in a partially cross-sectional approach.

Figure 2 displays cumulative birth-death actual values and predicted values for 5 selected approaches at the Total Private level from April 2019 - December 2020. The seasonal ARIMA model does not incorporate any concurrent information and predicts birth-death values consistent with historical trends-mostly positive. The modified version of RegARIMA models from Battista (2013) improves considerably in capturing some of the downturn in Mar.-Apr. 2020. The simple average of $\mathrm{S}_{2}$ provides substantially more gain, and the very simple partially-cross-sectional model (N11) and best ex post under M1 (N28) predictions track even closer to the actual birth-death values.

Figure 2: Selected Birth-Death Predictions and Actual Values, Total Private


## 5. Conclusion and future directions

In this paper, we showed that BLS can improve predictions of the actual birth-death values substantially beyond what has been done before using only the CES sample. The largest improvements were surrounding the COVID-19 pandemic recession, when BLS needed to make additional methodological changes to the estimator to reasonably capture business shutdowns and re-openings, but there were also substantial improvements surrounding the Great Recession and the ensuing recovery.

We found that how a model handles seasonality to be critical to its abilities to relate birth-death to the CES sample. Simple models that ignored seasonality and the time series nature of the data in a first step
fared quite well. Seasonal differencing and the use of seasonal dummies greatly attenuated the relationship between birth-death and the sample resulting in worse predictions during turning points.

Forecast combination provided good results, although considering the set of which forecasts to combine mattered. Some individual methods performed about as well across metrics as the combinations, although less consistently, and we intend to better understand why. Regardless we think that forecast combination will be useful to make predictions more robust to misspecification bias.

Future work should push to improve the overall birth-death error along the metrics outlined in this paper, but also address additional production and data user needs. We think that it is important to test the possible cointegrating relationship between the CES sample and birth-death and to investigate appropriate models that take this relationship into account. Recurrent Neural Networks remain a promising part of the effort to improve birth-death prediction and should be pursued further. The impact of prediction methods on various quality metrics (such as the basic level benchmark revisions) will need to be fully addressed. Finally, any change to birth-death prediction will need to consider its explainability to data users.

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| Rank | Method | Overall | 2008 | '09 | '10 | '11 | '12 | '13 | '14 | '15 | '16 | '17 | '18 | '19 | '20 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1 | N28 (Partially Cross-Sectional) | 10.7 | 4.7 | 6.8 | 9.3 | 8.8 | 8.1 | 5.2 | 4.6 | 6.7 | 7.3 | 6.1 | 5.6 | 5.5 | 30.7 |
| 2 | Simple Avg. (S2) | 10.9 | 4.7 | 6.1 | 6.6 | 6.8 | 8.1 | 4.6 | 4.4 | 6.2 | 7.1 | 5.7 | 5.4 | 5.1 | 33.2 |
| 3 | Bates-Granger (S2) | 10.9 | 4.7 | 6.2 | 6.7 | 7 | 8.1 | 4.7 | 4.4 | 6.2 | 7.2 | 5.8 | 5.4 | 5.1 | 33.4 |
| 4 | N31 (Partially Cross-Sectional) | 11 | 5 | 7.4 | 9.9 | 9.7 | 8.2 | 5.9 | 4.8 | 7.3 | 7.2 | 6.2 | 5.6 | 5.5 | 31.2 |
| 5 | N08 (Partially Cross-Sectional) | 11.1 | 4.7 | 6.6 | 6.2 | 6.8 | 8.1 | 4.5 | 4.5 | 6.4 | 7.2 | 6.3 | 5.3 | 5.5 | 34.1 |
| 6 | N11 (Partially Cross-Sectional) | 11.4 | 4.7 | 6.8 | 6.2 | 6.9 | 8.1 | 4.5 | 4.5 | 6.4 | 7.3 | 6.2 | 5.4 | 5.5 | 35.2 |
| 7 | N09 (Partially Cross-Sectional) | 11.8 | 4.7 | 6.4 | 6.1 | 6.7 | 8 | 4.5 | 4.5 | 6.3 | 7.2 | 6.1 | 5.4 | 5.5 | 37 |
| 8 | Bates-Granger (S1) | 12.5 | 5.2 | 6.2 | 6.1 | 6.3 | 8 | 4.4 | 4.3 | 6.2 | 7.1 | 5.7 | 5.2 | 5.1 | 40.2 |
| 9 | Simple Avg. (S1) | 12.8 | 5.2 | 6.1 | 6.1 | 6.2 | 8 | 4.4 | 4.3 | 6.2 | 7.1 | 5.7 | 5.1 | 5.1 | 41.3 |
| 10 | B08 (Partially Cross-Sectional) | 13.1 | 5.4 | 6.8 | 6 | 5.8 | 8.5 | 5 | 4.5 | 6.6 | 7.6 | 5.7 | 5.7 | 5.4 | 41.9 |
| 11 | B11 (Partially Cross-Sectional) | 13.1 | 5.1 | 6.7 | 6 | 5.9 | 8.6 | 5.1 | 4.5 | 6.6 | 7.6 | 5.6 | 5.7 | 5.5 | 41.9 |
| 12 | B60 (Partially Cross-Sectional) | 13.4 | 6.9 | 6.9 | 7.5 | 9 | 11 | 7.7 | 7.3 | 8.6 | 10 | 9.3 | 9.5 | 7.7 | 38.2 |
| 13 | Diebold-Shin (S2) | 13.4 | 4.7 | 6.2 | 6 | 6.3 | 8.3 | 4.8 | 4.5 | 6.3 | 7.1 | 5.7 | 5.2 | 5.1 | 43.8 |
| 14 | S35 (Partially Cross-Sectional) | 13.6 | 5.6 | 6.1 | 6.6 | 6.2 | 8.1 | 4.9 | 4.8 | 6 | 7 | 6.1 | 5.2 | 5.1 | 44.2 |
| 15 | S09 (Partially Cross-Sectional) | 13.6 | 5 | 6.4 | 6.3 | 6.9 | 8 | 4.9 | 4.4 | 6.2 | 7.3 | 5.9 | 5.3 | 5.3 | 44.2 |
| 16 | B09 (Partially Cross-Sectional) | 13.6 | 5.7 | 6.9 | 6 | 5.8 | 8.6 | 5 | 4.4 | 6.6 | 7.5 | 5.7 | 5.6 | 5.3 | 44.1 |
| 17 | S29 (Partially Cross-Sectional) | 13.8 | 5.4 | 6.4 | 9.3 | 8.4 | 8.1 | 5.1 | 4.5 | 6.2 | 7.3 | 5.9 | 5.4 | 5.3 | 44.4 |
| 18 | S08 (Partially Cross-Sectional) | 13.9 | 5.6 | 7.2 | 7.1 | 7.3 | 8.1 | 4.9 | 4.5 | 6.2 | 7.4 | 5.9 | 5.4 | 5.3 | 44.9 |
| 19 | S12 (Partially Cross-Sectional) | 14 | 6.2 | 6.6 | 6.1 | 6.6 | 8 | 4.6 | 4.4 | 6.3 | 7.4 | 6 | 5 | 5.3 | 45.9 |
| 20 | S28 (Partially Cross-Sectional) | 14.1 | 6 | 7.4 | 10.2 | 8.7 | 8.3 | 5.6 | 4.6 | 6.3 | 7.3 | 5.8 | 5.5 | 5.3 | 44.7 |
| 21 | S11 (Partially Cross-Sectional) | 14.4 | 5.9 | 8.2 | 7.8 | 7.5 | 8.1 | 5.1 | 4.6 | 6.3 | 7.4 | 5.9 | 5.4 | 5.3 | 46.8 |
| 22 | S15 (Partially Cross-Sectional) | 15 | 5.7 | 6.2 | 6.6 | 6.1 | 8.1 | 4.8 | 4.8 | 6.1 | 7.2 | 6.2 | 5.1 | 5.2 | 49.6 |
| 23 | B12 (Partially Cross-Sectional) | 15 | 5.9 | 7.2 | 5.8 | 5.8 | 8.4 | 4.7 | 4.7 | 6.7 | 7.3 | 5.7 | 5.4 | 5.1 | 49.8 |
| 24 | N12 (Partially Cross-Sectional) | 15 | 5.4 | 6.5 | 5.9 | 6 | 8 | 4.2 | 4.4 | 6.4 | 7.3 | 6 | 5 | 5.4 | 50.1 |
| 25 | Diebold-Shin (S1) | 15.4 | 5.2 | 6.1 | 6.4 | 6.5 | 8.3 | 4.6 | 4.4 | 6.4 | 7.1 | 5.7 | 5.1 | 5 | 51.6 |
| 26 | B29 (Partially Cross-Sectional) | 15.6 | 6.1 | 6.8 | 9.1 | 7.2 | 8.4 | 5.4 | 4.6 | 6.7 | 7.6 | 5.7 | 5.6 | 5.4 | 51.3 |
| 27 | S31 (Partially Cross-Sectional) | 15.9 | 6.4 | 7.9 | 11.1 | 10.9 | 8.3 | 6.3 | 4.8 | 6.8 | 7.4 | 6 | 5.6 | 5.4 | 51.3 |
| 28 | B28 (Partially Cross-Sectional) | 16.2 | 5.8 | 6.6 | 9.2 | 7.9 | 8.4 | 5.3 | 4.6 | 7 | 7.6 | 5.7 | 5.7 | 5.3 | 53.5 |
| 29 | S10 (Partially Cross-Sectional) | 16.9 | 6.9 | 7 | 6.1 | 6.4 | 7.9 | 4.6 | 4.5 | 6.4 | 7.3 | 5.9 | 4.9 | 5.2 | 57.1 |
| 30 | B10 (Partially Cross-Sectional) | 17.4 | 7 | 7.6 | 5.8 | 5.8 | 8.6 | 4.8 | 4.6 | 6.7 | 7.4 | 5.6 | 5.4 | 5.1 | 58.8 |
| 31 | N30 (Partially Cross-Sectional) | 17.4 | 6.5 | 7 | 8.6 | 6.3 | 7.9 | 4.3 | 4.5 | 6.4 | 7.3 | 6.2 | 5 | 5.3 | 58.7 |
| 32 | S27 (RegARIMA) | 17.5 | 6.3 | 6.4 | 6.3 | 6.6 | 8 | 4.6 | 4.5 | 6.4 | 7.2 | 5.9 | 4.9 | 5.1 | 59.6 |
| 33 | S30 (Partially Cross-Sectional) | 17.6 | 6.8 | 7 | 7 | 6.5 | 8.1 | 4.6 | 4.6 | 6.5 | 7.3 | 5.9 | 5 | 5.3 | 59.5 |
| 34 | S07 (RegARIMA) | 17.8 | 6.4 | 6.5 | 5.9 | 6.3 | 7.9 | 4.6 | 4.5 | 6.3 | 7.2 | 5.9 | 4.9 | 5.1 | 73.5 |
| 35 | N10 (Partially Cross-Sectional) | 17.8 | 6.5 | 7 | 6.1 | 5.9 | 8.2 | 4.2 | 4.5 | 6.3 | 7.1 | 6.1 | 5 | 5.4 | 60.7 |
| 36 | N35 (Partially Cross-Sectional) | 17.9 | 5.8 | 6.3 | 6.6 | 6.3 | 8.3 | 4.5 | 4.7 | 6.2 | 6.9 | 6.1 | 5.3 | 5.2 | 61 |
| 37 | B30 (Partially Cross-Sectional) | 18 | 6.9 | 7.6 | 8.7 | 7 | 8.5 | 4.4 | 4.6 | 6.7 | 7.4 | 5.7 | 5.4 | 5.2 | 60.7 |
| 38 | B31 (Partially Cross-Sectional) | 18.2 | 5.6 | 6.5 | 9.5 | 8.8 | 8.4 | 5.8 | 4.7 | 7.6 | 7.6 | 5.8 | 5.8 | 5.4 | 60.9 |
| 39 | B26 (RegARIMA) | 18.8 | 7.8 | 7.9 | 6.8 | 6.9 | 8.7 | 4.4 | 4.3 | 6.6 | 7.3 | 5.7 | 5.6 | 5.1 | 63.8 |
| 40 | N27 (RegARIMA) | 18.9 | 6.2 | 6.7 | 9.4 | 6.3 | 8 | 4.5 | 4.5 | 6.3 | 7.2 | 6.1 | 5 | 5.3 | 64.5 |


|  | S26 (RegARIMA) | 19.2 | 6.8 | 6.3 | 6.2 | 6.1 | 8.3 | 4.6 | 4.5 | 6.5 | 7 | 6.1 | 5 | 5.2 | 65.8 |
| :--- | :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| 42 | N07 (RegARIMA) | 19.4 | 6.3 | 6.9 | 5.9 | 5.7 | 8.1 | 4.3 | 4.4 | 6.3 | 7.2 | 6 | 5 | 5.3 | 81.1 |
| 43 | S06 (RegARIMA) | 19.5 | 6.2 | 6.5 | 5.6 | 6.2 | 7.9 | 4.7 | 4.5 | 6.5 | 7 | 6.1 | 4.9 | 5.1 | 67.1 |
| 44 | N26 (RegARIMA) | 19.6 | 6.7 | 6.7 | 9.5 | 6 | 8.1 | 4.5 | 4.6 | 6.2 | 7.2 | 6.3 | 4.9 | 5.3 | 67 |
| 45 | N06 (RegARIMA) | 19.6 | 6.7 | 6.9 | 5.8 | 5.7 | 8 | 4.4 | 4.5 | 6.3 | 7.2 | 6.1 | 4.9 | 5.4 | 67.5 |
| 46 | B07 (RegARIMA) | 19.7 | 6.5 | 7.6 | 6 | 5.6 | 8.6 | 4.7 | 4.5 | 6.7 | 7.3 | 5.8 | 5.5 | 5 | 67.7 |
| 47 | B27 (RegARIMA) | 19.8 | 7.5 | 7.7 | 6.7 | 6.5 | 8.6 | 4.5 | 4.4 | 6.6 | 7.2 | 5.7 | 5.6 | 5 | 67.7 |
| 48 | RegARIMA - Battista 2013 (modified) | 20.1 | 5.3 | 6.4 | 5.8 | 5.7 | 8.2 | 4.3 | 4.5 | 5.9 | 7.3 | 5.9 | 5.1 | 4.8 | 69.6 |
| 49 | B06 (RegARIMA) | 20.3 | 6.7 | 7.6 | 6 | 5.6 | 8.5 | 4.7 | 4.6 | 6.7 | 7.2 | 6 | 5.4 | 5.1 | 69.9 |
| 50 | N15 (Partially Cross-Sectional) | 20.4 | 5.7 | 6.3 | 6.4 | 5.8 | 8.2 | 4.4 | 4.7 | 6.1 | 7 | 6.1 | 5.1 | 5.3 | 70.5 |
| 51 | RegARIMA - Battista 2013 | 23.2 | 5.7 | 6.7 | 5.8 | 5.8 | 8.2 | 4.3 | 4.5 | 6.3 | 7.3 | 5.8 | 5.1 | 4.9 | 81.1 |
| 52 | DeepAR (RNN) | 25.1 | 8.6 | 10.4 | 7.9 | 6.9 | 8.1 | 6.8 | 5.6 | 7.5 | 7.8 | 7.7 | 7 | 8.9 | 86.4 |
| 53 | Production (ARIMA) | 27.1 | 6.9 | 7.7 | 6.2 | 5.4 | 8.2 | 4.3 | 4.4 | 6.2 | 7.2 | 6 | 4.9 | 5 | 95.3 |

## Appendix 2 - RMSE, Total Private Post-Benchmark 12-Mo. Summed Birth-Death, Thousands (M2)

| Rank | Method | Overall | 2008 | ‘09 | '10 | '11 | '12 | '13 | '14 | '15 | '16 | '17 | '18 | '19 | '20 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1 | B60 (Partially Cross-Sectional) | 156 | 125 | 105 | 19 | 38 | 278 | 229 | 97 | 61 | 138 | 25 | 83 | 243 | 246 |
| 2 | S09 (Partially Cross-Sectional) | 161 | 121 | 155 | 309 | 213 | 206 | 88 | 150 | 39 | 101 | 185 | 55 | 105 | 161 |
| 3 | S08 (Partially Cross-Sectional) | 167 | 98 | 43 | 347 | 291 | 182 | 87 | 141 | 29 | 101 | 189 | 61 | 102 | 152 |
| 4 | Diebold-Shin (S2) | 167 | 106 | 175 | 327 | 123 | 239 | 150 | 150 | 55 | 75 | 176 | 47 | 77 | 220 |
| 5 | S15 (Partially Cross-Sectional) | 169 | 142 | 167 | 328 | 117 | 244 | 106 | 135 | 22 | 34 | 113 | 9 | 160 | 267 |
| 6 | N11 (Partially Cross-Sectional) | 172 | 44 | 97 | 415 | 227 | 179 | 58 | 135 | 44 | 103 | 230 | 80 | 46 | 145 |
| 7 | Bates-Granger (S2) | 174 | 106 | 222 | 347 | 222 | 200 | 105 | 151 | 48 | 85 | 192 | 65 | 83 | 171 |
| 8 | Simple Avg. (S2) | 175 | 106 | 238 | 354 | 210 | 204 | 103 | 151 | 49 | 82 | 191 | 65 | 83 | 176 |
| 9 | N08 (Partially Cross-Sectional) | 176 | 80 | 181 | 409 | 181 | 203 | 85 | 147 | 47 | 99 | 221 | 72 | 45 | 164 |
| 10 | S35 (Partially Cross-Sectional) | 177 | 225 | 145 | 340 | 124 | 248 | 105 | 142 | 15 | 31 | 121 | 9 | 160 | 263 |
| 11 | Diebold-Shin (S1) | 179 | 142 | 226 | 339 | 114 | 259 | 130 | 165 | 68 | 90 | 185 | 39 | 73 | 233 |
| 12 | S11 (Partially Cross-Sectional) | 180 | 53 | 116 | 369 | 369 | 134 | 72 | 152 | 24 | 110 | 187 | 62 | 104 | 148 |
| 13 | N09 (Partially Cross-Sectional) | 182 | 82 | 248 | 403 | 178 | 211 | 97 | 155 | 57 | 96 | 207 | 80 | 49 | 166 |
| 14 | S29 (Partially Cross-Sectional) | 184 | 143 | 192 | 306 | 357 | 194 | 61 | 165 | 46 | 88 | 196 | 64 | 107 | 166 |
| 15 | S28 (Partially Cross-Sectional) | 186 | 105 | 87 | 337 | 411 | 178 | 43 | 144 | 43 | 89 | 213 | 66 | 104 | 154 |
| 16 | N31 (Partially Cross-Sectional) | 186 | 91 | 113 | 465 | 297 | 119 | 18 | 102 | 64 | 103 | 244 | 83 | 48 | 130 |
| 17 | N28 (Partially Cross-Sectional) | 187 | 126 | 214 | 426 | 261 | 152 | 39 | 148 | 71 | 94 | 224 | 72 | 52 | 151 |
| 18 | Bates-Granger (S1) | 199 | 142 | 395 | 342 | 131 | 250 | 124 | 170 | 66 | 98 | 189 | 59 | 73 | 201 |
| 19 | N15 (Partially Cross-Sectional) | 200 | 145 | 416 | 271 | 38 | 316 | 139 | 145 | 5 | 24 | 100 | 10 | 133 | 283 |
| 20 | S12 (Partially Cross-Sectional) | 200 | 178 | 421 | 298 | 76 | 271 | 109 | 174 | 57 | 87 | 173 | 19 | 105 | 218 |
| 21 | Simple Avg. (S1) | 201 | 142 | 398 | 348 | 129 | 252 | 123 | 171 | 67 | 95 | 188 | 58 | 73 | 206 |
| 22 | S31 (Partially Cross-Sectional) | 204 | 43 | 43 | 359 | 537 | 94 | 17 | 142 | 49 | 88 | 226 | 62 | 99 | 127 |
| 23 | S06 (RegARIMA) | 205 | 185 | 454 | 275 | 137 | 253 | 91 | 164 | 41 | 85 | 176 | 19 | 94 | 249 |
| 24 | S26 (RegARIMA) | 209 | 317 | 415 | 270 | 141 | 240 | 49 | 174 | 45 | 86 | 182 | 13 | 90 | 250 |
| 25 | B31 (Partially Cross-Sectional) | 211 | 44 | 434 | 393 | 222 | 160 | 138 | 151 | 102 | 122 | 240 | 133 | 14 | 124 |


| 26 | B11 (Partially Cross-Sectional) | 211 | 57 | 408 | 386 | 179 | 216 | 158 | 193 | 105 | 124 | 235 | 139 | 22 | 133 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 27 | N35 (Partially Cross-Sectional) | 213 | 261 | 462 | 281 | 5 | 282 | 135 | 149 | 14 | 17 | 108 | 4 | 133 | 280 |
| 28 | RegARIMA - Battista 2013 (modified) | 214 | 76 | 416 | 373 | 136 | 301 | 187 | 209 | 62 | 101 | 180 | 63 | 56 | 197 |
| 29 | B28 (Partially Cross-Sectional) | 219 | 50 | 479 | 389 | 183 | 193 | 142 | 166 | 105 | 124 | 240 | 125 | 22 | 146 |
| 30 | N12 (Partially Cross-Sectional) | 220 | 148 | 482 | 352 | 77 | 274 | 130 | 185 | 76 | 89 | 202 | 57 | 59 | 235 |
| 31 | B08 (Partially Cross-Sectional) | 223 | 76 | 459 | 388 | 158 | 239 | 173 | 198 | 109 | 134 | 239 | 134 | 26 | 146 |
| 32 | S07 (RegARIMA) | 223 | 181 | 521 | 325 | 28 | 304 | 111 | 170 | 60 | 87 | 166 | 20 | 98 | 237 |
| 33 | RegARIMA - Battista 2013 | 225 | 104 | 486 | 355 | 150 | 297 | 176 | 203 | 72 | 101 | 168 | 55 | 54 | 225 |
| 34 | S27 (RegARIMA) | 226 | 244 | 494 | 326 | 69 | 324 | 107 | 167 | 57 | 86 | 166 | 17 | 98 | 239 |
| 35 | N26 (RegARIMA) | 227 | 261 | 542 | 316 | 43 | 252 | 72 | 188 | 60 | 79 | 164 | 45 | 60 | 250 |
| 36 | B09 (Partially Cross-Sectional) | 227 | 86 | 484 | 386 | 130 | 248 | 184 | 203 | 115 | 134 | 237 | 119 | 28 | 162 |
| 37 | B29 (Partially Cross-Sectional) | 229 | 92 | 516 | 396 | 151 | 227 | 149 | 165 | 109 | 124 | 229 | 121 | 21 | 169 |
| 38 | N06 (RegARIMA) | 231 | 177 | 566 | 305 | 6 | 299 | 120 | 192 | 52 | 81 | 167 | 48 | 69 | 252 |
| 39 | S10 (Partially Cross-Sectional) | 233 | 205 | 563 | 310 | 58 | 303 | 116 | 177 | 67 | 92 | 164 | 8 | 103 | 243 |
| 40 | N07 (RegARIMA) | 236 | 182 | 564 | 321 | 62 | 321 | 149 | 191 | 62 | 74 | 163 | 47 | 70 | 253 |
| 41 | S30 (Partially Cross-Sectional) | 240 | 228 | 557 | 324 | 134 | 333 | 115 | 172 | 63 | 92 | 162 | 10 | 101 | 243 |
| 42 | N27 (RegARIMA) | 241 | 197 | 575 | 343 | 1 | 333 | 117 | 191 | 72 | 72 | 158 | 42 | 68 | 251 |
| 43 | N10 (Partially Cross-Sectional) | 245 | 194 | 596 | 346 | 29 | 315 | 148 | 189 | 68 | 86 | 185 | 52 | 58 | 246 |
| 44 | N30 (Partially Cross-Sectional) | 249 | 208 | 589 | 358 | 93 | 325 | 129 | 194 | 80 | 89 | 185 | 49 | 60 | 252 |
| 45 | DeepAR (RNN) | 253 | 83 | 713 | 292 | 56 | 82 | 184 | 105 | 56 | 5 | 147 | 111 | 143 | 343 |
| 46 | B12 (Partially Cross-Sectional) | 253 | 97 | 581 | 406 | 56 | 301 | 201 | 209 | 135 | 159 | 217 | 96 | 29 | 206 |
| 47 | B26 (RegARIMA) | 263 | 172 | 670 | 352 | 59 | 304 | 162 | 200 | 103 | 153 | 201 | 80 | 19 | 228 |
| 48 | B06 (RegARIMA) | 265 | 150 | 654 | 352 | 53 | 311 | 211 | 212 | 143 | 157 | 215 | 77 | 21 | 230 |
| 49 | B07 (RegARIMA) | 275 | 129 | 681 | 367 | 31 | 343 | 224 | 203 | 156 | 164 | 209 | 92 | 29 | 235 |
| 50 | B30 (Partially Cross-Sectional) | 277 | 119 | 699 | 413 | 29 | 311 | 178 | 214 | 131 | 155 | 207 | 101 | 25 | 230 |
| 51 | B10 (Partially Cross-Sectional) | 277 | 129 | 694 | 397 | 15 | 316 | 212 | 218 | 158 | 155 | 207 | 92 | 31 | 220 |
| 52 | B27 (RegARIMA) | 279 | 201 | 685 | 389 | 70 | 326 | 214 | 209 | 131 | 164 | 209 | 92 | 29 | 235 |
| 53 | Production (ARIMA) | 293 | 179 | 723 | 345 | 157 | 433 | 251 | 222 | 108 | 115 | 149 | 38 | 65 | 257 |

Appendix 3 - RMSE, Total Private 12-Mo. Summed Birth-Death from December, Thousands (M2 - alternate window)

| Rank | Method | Overall | 2008 | '09 | '10 | '11 | '12 | '13 | '14 | '15 | '16 | '17 | '18 | '19 | '20 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1 | B60 (Partially Cross-Sectional) | 147 | 93 | 13 | 15 | 41 | 459 | 52 | 45 | 1 | 124 | 83 | 146 | 98 | 56 |
| 2 | S15 (Partially Cross-Sectional) | 150 | 181 | 217 | 259 | 11 | 332 | 10 | 14 | 162 | 52 | 3 | 41 | 70 | 13 |
| 3 | S35 (Partially Cross-Sectional) | 154 | 164 | 227 | 263 | 13 | 336 | 8 | 24 | 162 | 57 | 9 | 44 | 71 | 100 |
| 4 | N11 (Partially Cross-Sectional) | 156 | 125 | 310 | 234 | 73 | 278 | 12 | 44 | 212 | 73 | 98 | 56 | 21 | 41 |
| 5 | S09 (Partially Cross-Sectional) | 156 | 178 | 271 | 222 | 56 | 317 | 22 | 46 | 229 | 38 | 52 | 23 | 9 | 16 |
| 6 | S08 (Partially Cross-Sectional) | 160 | 106 | 237 | 295 | 72 | 323 | 6 | 40 | 231 | 46 | 56 | 22 | 3 | 85 |
| 7 | N08 (Partially Cross-Sectional) | 165 | 179 | 346 | 197 | 45 | 306 | 27 | 52 | 213 | 64 | 95 | 55 | 17 | 45 |
| 8 | Diebold-Shin (S2) | 169 | 220 | 260 | 176 | 4 | 353 | 45 | 61 | 196 | 13 | 67 | 25 | 2 | 223 |
| 9 | Bates-Granger (S2) | 169 | 220 | 303 | 244 | 51 | 336 | 31 | 57 | 205 | 37 | 74 | 28 | 6 | 51 |
| 10 | Simple Avg. (S2) | 170 | 220 | 317 | 234 | 46 | 334 | 33 | 57 | 203 | 35 | 73 | 28 | 3 | 72 |
| 11 | S11 (Partially Cross-Sectional) | 171 | 14 | 168 | 382 | 129 | 303 | 8 | 32 | 236 | 53 | 54 | 20 | 5 | 179 |
| 12 | S29 (Partially Cross-Sectional) | 172 | 174 | 259 | 362 | 80 | 298 | 13 | 54 | 219 | 48 | 67 | 17 | 6 | 30 |
| 13 | N09 (Partially Cross-Sectional) | 173 | 225 | 360 | 188 | 44 | 307 | 36 | 63 | 217 | 51 | 92 | 56 | 15 | 97 |
| 14 | S28 (Partially Cross-Sectional) | 176 | 117 | 227 | 407 | 94 | 299 | 39 | 49 | 227 | 56 | 81 | 9 | 5 | 87 |
| 15 | N28 (Partially Cross-Sectional) | 181 | 220 | 368 | 282 | 86 | 282 | 20 | 68 | 226 | 73 | 90 | 55 | 31 | 4 |
| 16 | N31 (Partially Cross-Sectional) | 183 | 166 | 328 | 354 | 107 | 266 | 66 | 52 | 223 | 72 | 107 | 53 | 30 | 128 |
| 17 | N35 (Partially Cross-Sectional) | 195 | 389 | 342 | 69 | 57 | 379 | 34 | 44 | 155 | 49 | 22 | 28 | 74 | 195 |
| 18 | N15 (Partially Cross-Sectional) | 204 | 371 | 328 | 23 | 77 | 375 | 48 | 24 | 152 | 57 | 10 | 22 | 71 | 335 |
| 19 | S31 (Partially Cross-Sectional) | 206 | 24 | 155 | 535 | 193 | 275 | 66 | 37 | 228 | 74 | 90 | 9 | 22 | 240 |
| 20 | S12 (Partially Cross-Sectional) | 213 | 369 | 364 | 85 | 23 | 350 | 50 | 84 | 212 | 1 | 42 | 17 | 25 | 368 |
| 21 | Bates-Granger (S1) | 214 | 347 | 377 | 156 | 3 | 362 | 57 | 82 | 214 | 28 | 75 | 37 | 9 | 334 |
| 22 | Simple Avg. (S1) | 217 | 347 | 380 | 155 | 2 | 361 | 58 | 82 | 212 | 26 | 74 | 37 | 6 | 359 |
| 23 | B11 (Partially Cross-Sectional) | 221 | 322 | 452 | 147 | 35 | 372 | 107 | 103 | 223 | 80 | 125 | 89 | 53 | 245 |
| 24 | Diebold-Shin (S1) | 222 | 347 | 281 | 208 | 21 | 357 | 37 | 81 | 216 | 28 | 55 | 24 | 9 | 463 |
| 25 | S26 (RegARIMA) | 224 | 416 | 301 | 169 | 20 | 304 | 20 | 96 | 207 | 12 | 46 | 18 | 7 | 461 |
| 26 | B31 (Partially Cross-Sectional) | 225 | 314 | 448 | 265 | 63 | 364 | 89 | 95 | 230 | 72 | 125 | 88 | 63 | 225 |
| 27 | S06 (RegARIMA) | 227 | 433 | 299 | 141 | 8 | 335 | 29 | 84 | 208 | 2 | 43 | 17 | 11 | 454 |
| 28 | B08 (Partially Cross-Sectional) | 236 | 357 | 472 | 125 | 33 | 386 | 125 | 105 | 225 | 82 | 130 | 90 | 55 | 314 |
| 29 | B28 (Partially Cross-Sectional) | 237 | 344 | 469 | 217 | 34 | 374 | 103 | 100 | 233 | 78 | 130 | 84 | 50 | 311 |
| 30 | B09 (Partially Cross-Sectional) | 245 | 380 | 492 | 103 | 19 | 391 | 124 | 114 | 230 | 79 | 131 | 79 | 57 | 339 |
| 31 | B29 (Partially Cross-Sectional) | 248 | 386 | 503 | 193 | 5 | 384 | 103 | 111 | 229 | 71 | 124 | 86 | 46 | 327 |
| 32 | S07 (RegARIMA) | 253 | 446 | 403 | 64 | 51 | 365 | 46 | 85 | 215 | 5 | 41 | 16 | 16 | 524 |
| 33 | S27 (RegARIMA) | 254 | 416 | 382 | 111 | 80 | 362 | 39 | 81 | 218 | 6 | 43 | 15 | 15 | 558 |
| 34 | N12 (Partially Cross-Sectional) | 258 | 402 | 419 | 93 | 16 | 367 | 77 | 85 | 216 | 28 | 86 | 46 | 8 | 562 |
| 35 | N06 (RegARIMA) | 274 | 514 | 374 | 12 | 0 | 391 | 63 | 86 | 198 | 17 | 82 | 47 | 2 | 597 |
| 36 | S30 (Partially Cross-Sectional) | 274 | 458 | 401 | 180 | 84 | 364 | 52 | 88 | 212 | 2 | 33 | 14 | 19 | 616 |
| 37 | S10 (Partially Cross-Sectional) | 275 | 491 | 419 | 84 | 48 | 365 | 57 | 93 | 217 | 10 | 36 | 14 | 23 | 606 |
| 38 | N26 (RegARIMA) | 278 | 514 | 335 | 72 | 11 | 360 | 28 | 100 | 197 | 14 | 66 | 49 | 1 | 663 |
| 39 | RegARIMA - Battista 2013 | 283 | 410 | 338 | 159 | 2 | 399 | 108 | 92 | 212 | 25 | 60 | 70 | 52 | 703 |
| 40 | N07 (RegARIMA) | 290 | 499 | 454 | 9 | 43 | 408 | 92 | 82 | 204 | 21 | 74 | 53 | 5 | 638 |


| 41 | N27 (RegARIMA) | 292 | 504 | 426 | 79 | 101 | 388 | 65 | 88 | 209 | 24 | 65 | 44 | 5 | 664 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 42 | N10 (Partially Cross-Sectional) | 299 | 515 | 449 | 65 | 42 | 395 | 93 | 87 | 211 | 7 | 79 | 45 | 5 | 681 |
| 43 | N30 (Partially Cross-Sectional) | 299 | 525 | 438 | 167 | 94 | 383 | 75 | 97 | 213 | 18 | 70 | 48 | 5 | 668 |
| 44 | B12 (Partially Cross-Sectional) | 303 | 433 | 558 | 64 | 33 | 428 | 140 | 158 | 262 | 70 | 107 | 77 | 49 | 607 |
| 45 | DeepAR (RNN) <br> RegARIMA - Battista 2013 | 311 | 572 | 622 | 131 | 61 | 232 | 215 | 97 | 196 | 155 | 35 | 33 | 65 | 590 |
| 46 | (modified) | 323 | 328 | 360 | 147 | 16 | 408 | 113 | 93 | 213 | 40 | 66 | 60 | 44 | 924 |
| 47 | B06 (RegARIMA) | 337 | 516 | 558 | 25 | 24 | 435 | 132 | 158 | 240 | 56 | 112 | 72 | 45 | 765 |
| 48 | B10 (Partially Cross-Sectional) | 339 | 523 | 591 | 24 | 52 | 445 | 156 | 163 | 261 | 60 | 116 | 69 | 45 | 725 |
| 49 | B30 (Partially Cross-Sectional) | 339 | 517 | 580 | 100 | 96 | 430 | 139 | 148 | 243 | 66 | 113 | 75 | 31 | 752 |
| 50 | B26 (RegARIMA) | 343 | 581 | 525 | 14 | 79 | 404 | 120 | 128 | 233 | 45 | 97 | 62 | 50 | 801 |
| 51 | B07 (RegARIMA) | 345 | 535 | 578 | 16 | 55 | 458 | 144 | 142 | 249 | 57 | 116 | 72 | 40 | 769 |
| 52 | B27 (RegARIMA) | 351 | 608 | 579 | 36 | 89 | 432 | 136 | 143 | 244 | 54 | 95 | 74 | 41 | 765 |
| 53 | Production (ARIMA) | 384 | 558 | 587 | 200 | 177 | 492 | 142 | 135 | 226 | 13 | 72 | 61 | 7 | 924 |

Appendix 4 - RMSE, Super Sector Post-Benchmark 12-Mo. Summed Birth-Death, Thousands

| Rank | Method | Overall | 2008 | ‘09 | '10 | '11 | '12 | '13 | '14 | '15 | '16 | '17 | '18 | '19 | '20 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1 | Diebold-Shin (S2) | 24.3 | 21.8 | 33.8 | 39.2 | 16.8 | 26.4 | 17.5 | 13.6 | 7.7 | 16.9 | 22.2 | 16.2 | 20.5 | 39.4 |
| 2 | Simple Avg. (S2) | 25 | 21.8 | 37.8 | 39.9 | 22.9 | 27.1 | 13.8 | 13.9 | 7.8 | 19 | 28.5 | 17 | 20.3 | 32.1 |
| 3 | Diebold-Shin (S1) | 25.3 | 22.2 | 38.6 | 39.9 | 15.7 | 27.9 | 16.3 | 14.9 | 8.7 | 18.4 | 23.4 | 15.8 | 20.3 | 40 |
| 4 | Bates-Granger (S2) | 25.4 | 21.8 | 40.5 | 39.6 | 23.9 | 27 | 13.9 | 13.7 | 7.8 | 19.5 | 28.7 | 17.3 | 20.6 | 32 |
| 5 | S09 (Partially Cross-Sectional) | 25.6 | 21.8 | 38.7 | 36.1 | 30.9 | 31.6 | 14.2 | 14.9 | 7.9 | 20.2 | 31.9 | 15.9 | 19.3 | 27.3 |
| 6 | S15 (Partially Cross-Sectional) | 25.8 | 22 | 37.7 | 38.6 | 23.3 | 30.8 | 16.3 | 13.6 | 8.1 | 14.9 | 21.5 | 15.4 | 23.4 | 42.5 |
| 7 | Simple Avg. (S1) | 26.1 | 22.2 | 46.4 | 38.4 | 17.6 | 28.4 | 14.3 | 15.4 | 8.6 | 19.6 | 27.1 | 16.8 | 19.8 | 36.5 |
| 8 | Bates-Granger (S1) <br> RegARIMA - Battista 2013 | 26.3 | 22.2 | 47.7 | 38 | 17.3 | 28.4 | 14.3 | 15.3 | 8.5 | 19.8 | 27.2 | 17 | 20 | 36.2 |
| 9 | (modified) | 26.6 | 15.6 | 45.3 | 42 | 19.7 | 29.3 | 20.3 | 19.2 | 8.2 | 21.8 | 23.6 | 16.6 | 16.8 | 38.6 |
| 10 | B31 (Partially Cross-Sectional) | 26.8 | 19.5 | 49 | 40.9 | 23.5 | 28 | 14.2 | 13.7 | 12.1 | 22.2 | 31.2 | 21 | 19.6 | 26.7 |
| 11 | S35 (Partially Cross-Sectional) | 26.8 | 36.1 | 36.7 | 38.8 | 23.7 | 30.5 | 16.8 | 13.7 | 7.9 | 15 | 21.9 | 15.2 | 23.5 | 41.2 |
| 12 | N09 (Partially Cross-Sectional) | 27 | 22.7 | 36.3 | 47.9 | 25.7 | 30.9 | 14.8 | 14.9 | 9.7 | 20.5 | 31.4 | 17.8 | 20 | 33.3 |
| 13 | N15 (Partially Cross-Sectional) | 27.2 | 25 | 51.1 | 33.3 | 20 | 33.8 | 15.4 | 14 | 6.7 | 13 | 16.9 | 17.9 | 22.7 | 44.4 |
| 14 | B28 (Partially Cross-Sectional) | 27.5 | 20.7 | 52.6 | 40.9 | 20.7 | 28.5 | 14 | 14.9 | 11.9 | 21.8 | 31.5 | 20.4 | 20.8 | 28.8 |
| 15 | S29 (Partially Cross-Sectional) | 27.5 | 31.3 | 39.5 | 36.1 | 41.6 | 28.9 | 14.8 | 15.2 | 7.9 | 20.1 | 32 | 16.7 | 19.5 | 28.5 |
| 16 | B11 (Partially Cross-Sectional) | 27.5 | 19.9 | 47.8 | 41.4 | 24.3 | 25.8 | 18.1 | 18.6 | 11.7 | 22.9 | 31.7 | 20.9 | 19.2 | 32.6 |
| 17 | N08 (Partially Cross-Sectional) | 27.7 | 25.6 | 35.4 | 50.1 | 25.5 | 32.1 | 15.4 | 14.6 | 9.8 | 20.9 | 32.7 | 18 | 19.9 | 33.5 |
| 18 | N11 (Partially Cross-Sectional) | 28 | 23.5 | 35.6 | 51.7 | 28.6 | 32 | 16.3 | 13.9 | 10.1 | 21.6 | 32.8 | 17.9 | 19.7 | 32.5 |
| 19 | RegARIMA - Battista 2013 | 28.1 | 16.8 | 54.1 | 41 | 21.4 | 29.2 | 19.4 | 18.7 | 9 | 21.7 | 22.2 | 16.3 | 16.9 | 41.9 |
| 20 | S06 (RegARIMA) | 28.2 | 23.5 | 56.8 | 31.7 | 29.8 | 30.5 | 13.5 | 15.3 | 8.2 | 18.5 | 27.5 | 14.9 | 18.7 | 39.5 |
| 21 | S12 (Partially Cross-Sectional) | 28.2 | 27.2 | 55.3 | 32.7 | 29.6 | 33.5 | 14.9 | 16 | 9.8 | 18.7 | 28.7 | 15.1 | 19.8 | 33.2 |
| 22 | N28 (Partially Cross-Sectional) | 28.2 | 28.2 | 34.9 | 49.8 | 31.3 | 32.6 | 13.5 | 14.4 | 11.1 | 20.5 | 33.2 | 17.5 | 20.8 | 32.1 |
| 23 | B29 (Partially Cross-Sectional) | 28.2 | 20.4 | 55.4 | 41.9 | 19.6 | 28.1 | 14.6 | 15 | 11.8 | 22.2 | 30.7 | 19.9 | 20.7 | 33.3 |
| 24 | B08 (Partially Cross-Sectional) | 28.2 | 19.2 | 51.8 | 42.1 | 23.6 | 26.4 | 19.8 | 19 | 11.9 | 22.7 | 32.2 | 20.4 | 19.8 | 32.2 |
| 25 | S26 (RegARIMA) | 28.5 | 37.5 | 51.7 | 31.7 | 30.8 | 27 | 15.1 | 15.8 | 6.9 | 17.7 | 27.6 | 15.4 | 18.7 | 40 |


| 26 | S27 (RegARIMA) | 28.6 | 29.5 | 54.4 | 36.5 | 23.4 | 32.8 | 14.7 | 15.8 | 10.2 | 19.6 | 27 | 14.8 | 18.7 | 39.7 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 27 | S07 (RegARIMA) | 28.6 | 25.2 | 57.6 | 36 | 24.1 | 32.4 | 15.4 | 16 | 10.2 | 19 | 26.9 | 14.8 | 18.8 | 39 |
| 28 | N35 (Partially Cross-Sectional) | 28.8 | 41.2 | 54.1 | 31 | 21 | 31.1 | 16.7 | 14.2 | 6.1 | 13.8 | 19.4 | 16.3 | 23.3 | 43.1 |
| 29 | B09 (Partially Cross-Sectional) | 28.8 | 19.9 | 53.3 | 42.1 | 22.3 | 27.5 | 20.4 | 19.7 | 12.2 | 22.8 | 31.9 | 19.8 | 19.9 | 34.9 |
| 30 | N31 (Partially Cross-Sectional) | 28.9 | 26.7 | 36.2 | 53 | 32.9 | 32.8 | 15.2 | 13.3 | 11 | 21.7 | 34.4 | 17.7 | 20.1 | 31.6 |
| 31 | N12 (Partially Cross-Sectional) | 29 | 24.1 | 55.6 | 40.1 | 21.6 | 33.2 | 15 | 17.2 | 11.1 | 19.1 | 28.2 | 17.2 | 20.2 | 39.7 |
| 32 | N26 (RegARIMA) | 29.1 | 32.4 | 60.5 | 35.7 | 18.3 | 26 | 13.4 | 17.8 | 7.5 | 17.8 | 24.7 | 18 | 19.3 | 43.1 |
| 33 | S08 (Partially Cross-Sectional) | 29.1 | 24.9 | 50.5 | 44.2 | 41.3 | 29.4 | 15 | 14.1 | 7.9 | 20.7 | 32.3 | 16.3 | 18.9 | 27.6 |
| 34 | N07 (RegARIMA) | 29.6 | 26.9 | 62.5 | 36.6 | 15.6 | 32.1 | 16.2 | 17.7 | 8.6 | 17.4 | 23.1 | 18.2 | 20.8 | 43.3 |
| 35 | N06 (RegARIMA) | 29.8 | 24.7 | 66 | 34.7 | 16 | 30.5 | 14.2 | 18.4 | 7.6 | 18 | 24.7 | 18.2 | 20.3 | 43.4 |
| 36 | N27 (RegARIMA) | 29.9 | 26.6 | 63.6 | 37.7 | 19.1 | 31.6 | 14.1 | 17.8 | 9.2 | 17.8 | 22.9 | 18.2 | 20.9 | 43 |
| 37 | S28 (Partially Cross-Sectional) | 29.9 | 30 | 50.7 | 40.8 | 47.3 | 27.6 | 16.9 | 13.5 | 7.9 | 20.2 | 32.9 | 17.2 | 19 | 28.5 |
| 38 | S10 (Partially Cross-Sectional) | 30.6 | 30.5 | 66.5 | 34.2 | 28.3 | 33.6 | 15 | 16.5 | 10.6 | 19 | 26.7 | 15.3 | 19.6 | 37.6 |
| 39 | B12 (Partially Cross-Sectional) | 31 | 22 | 63 | 43.7 | 20.1 | 30.5 | 20.5 | 19.2 | 14.6 | 25.7 | 27.6 | 18.8 | 20 | 40.8 |
| 40 | N10 (Partially Cross-Sectional) | 31.1 | 27.7 | 67.8 | 39.6 | 20.4 | 33.8 | 16 | 17.7 | 9 | 18.7 | 25 | 17.3 | 20.6 | 41.5 |
| 41 | S30 (Partially Cross-Sectional) | 31.1 | 31.8 | 66.9 | 35.5 | 29.9 | 34.7 | 13.7 | 16 | 10.3 | 19.4 | 26.4 | 15.2 | 19.5 | 38.5 |
| 42 | S11 (Partially Cross-Sectional) | 31.4 | 22.3 | 59.5 | 47.6 | 48 | 27.3 | 15.5 | 15 | 7.8 | 21.4 | 32.2 | 16.6 | 18.8 | 27.9 |
| 43 | N30 (Partially Cross-Sectional) | 31.5 | 28.5 | 67.5 | 40.7 | 23.8 | 33.5 | 14.4 | 17.6 | 9.6 | 18.9 | 27 | 17.2 | 20.3 | 42.2 |
| 44 | S31 (Partially Cross-Sectional) | 31.5 | 29.6 | 52.5 | 41.5 | 56.6 | 26.2 | 17.9 | 13.9 | 9.3 | 22.9 | 33.8 | 18 | 18.6 | 26.8 |
| 45 | B26 (RegARIMA) | 32.2 | 28.7 | 71.2 | 37.6 | 19.5 | 33.2 | 14.8 | 18.7 | 11.4 | 24.1 | 26.5 | 19 | 20.7 | 44.2 |
| 46 | B06 (RegARIMA) | 32.3 | 24.9 | 69.6 | 39.7 | 20.7 | 30.6 | 21.3 | 19.8 | 15.6 | 24.3 | 27.2 | 18.3 | 20.8 | 44.6 |
| 47 | B60 (Partially Cross-Sectional) | 32.5 | 32.8 | 45 | 27.9 | 28.1 | 25.4 | 23.6 | 20.2 | 24.5 | 34 | 26.9 | 38.7 | 42.3 | 40.7 |
| 48 | B07 (RegARIMA) | 33 | 24.8 | 72.2 | 40.9 | 18 | 33.9 | 22 | 18.2 | 16.1 | 25.3 | 26.1 | 19.3 | 20.2 | 44.4 |
| 49 | B30 (Partially Cross-Sectional) | 33.1 | 24.2 | 75.2 | 44.9 | 15.3 | 30.9 | 16.6 | 19.4 | 14.3 | 25.2 | 26.4 | 19.5 | 20.6 | 43 |
| 50 | B10 (Partially Cross-Sectional) | 33.1 | 23.9 | 73.3 | 43.7 | 20.8 | 31.2 | 21 | 19.8 | 16.8 | 25.2 | 26.5 | 18.6 | 19 | 43.2 |
| 51 | B27 (RegARIMA) | 33.3 | 30.7 | 73.5 | 41.2 | 18.5 | 32 | 20.5 | 18.8 | 13.8 | 25.3 | 26.1 | 19.3 | 20.2 | 44.4 |
| 52 | Production (ARIMA) | 33.8 | 27.3 | 75.8 | 37.6 | 26.5 | 40.6 | 25.1 | 20.4 | 12.1 | 23.2 | 19.7 | 16 | 17.2 | 44 |
| 53 | DeepAR (RNN) | 34.4 | 30.3 | 77.8 | 38.9 | 23.7 | 16.6 | 24.4 | 15.5 | 17 | 22.3 | 32.4 | 18.3 | 21.7 | 51 |


[^0]:    ${ }^{1}$ Any opinions expressed in this paper are those of the authors and do not constitute policy of the Bureau of Labor Statistics.

[^1]:    ${ }^{2}$ Basic estimating cells comprise an industry or set of industries defined by the North American Industry Classification (NAICS) code, often at the 6-digit NAICS level of detail, sometimes broken out by Census region.

[^2]:    ${ }^{3} B^{n} \times X_{t}=X_{t-n}$
    ${ }^{4}$ Seasonal ARIMA models of the type examined in this paper are explained in the Reference Manual for X13-ARIMA-SEATS, Chapter 4: RegARIMA Modelling Capabilities, available at: https://www2.census.gov/software/x-13arima-seats/x-13-data/documentation/docx13as.pdf

[^3]:    ${ }^{5}$ The exception is when additive (point) outliers are included in the $X$ matrix, but these are not known for the forecast period. Additive outliers are also described in the X-13ARIMA-SEATS chapter on RegARIMA.
    ${ }^{6}$ In some models we do allow for nonlinear relationships in which $B D_{t}=f\left(X_{t}\right)+z_{t}$.

[^4]:    ${ }^{7}$ Super sectors are high-level industry combinations, mostly defined at the 2-digit NAICS level of detail, such as construction, retail trade, and leisure and hospitality.
    ${ }^{8}$ This was done by fitting conditional inference trees in the R package party (Hothorn et al. 2023).

[^5]:    ${ }^{9}$ With 44 forecasts and $K=5$ this requires 1,086,008 combinations evaluated for each super sector, each quarter.

